Parametric Signal Modeling and Linear Prediction Theory 5. Lattice Predictor

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5 Lattice Predictor

Appendix: Detailed Derivations

5.1 Basic Lattice Structure 5.2 Correlation Properties

5.2 Correlation Properties5.3 Joint Process Estimator

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Introduction

Recall: a forward or backward prediction-error filter can each be realized using a separate tapped-delay-line structure.

Lattice structure discussed in this section provides a powerful way to combine the FLP and BLP operations into a **single** structure.

Order Update for Prediction Errors

(Readings: Haykin §3.8)

Review:

Levinson-Durbin recursion:

$$\underline{a}_m = \left[\begin{array}{c} \underline{a}_{m-1} \\ 0 \end{array} \right] + \Gamma_m \left[\begin{array}{c} 0 \\ \underline{a}_{m-1}^{B^*} \end{array} \right]$$
 (forward)

$$\underline{a}_{m}^{B^{*}} = \begin{bmatrix} 0 \\ \underline{a}_{m-1}^{B^{*}} \end{bmatrix} + \Gamma_{m}^{*} \begin{bmatrix} \underline{a}_{m-1} \\ 0 \end{bmatrix}$$
 (backward)

Recursive Relations for $f_m[n]$ and $b_m[n]$

$$f_m[n] = \underline{a}_m^H \underline{u}_{m+1}[n]; \ b_m[n] = \underline{a}_m^{BT} \underline{u}_{m+1}[n]$$

• FLP: $f_m[n] = \left[\underline{a}_{m-1}^H \mid 0 \right] \left[\begin{array}{c} \underline{u}_m[n] \\ u[n-m] \end{array} \right] + \Gamma_m^* \left[\begin{array}{c} 0 \mid \underline{a}_{m-1}^{BT} \end{array} \right] \left[\begin{array}{c} u[n] \\ \underline{u}_m[n-1] \end{array} \right]$ (Details)

$$f_m[n] = f_{m-1}[n] + \Gamma_m^* b_{m-1}[n-1]$$

3 BLP: $b_m[n] = \left[0 \middle| \underline{a}_{m-1}^{BT} \right] \left[\begin{array}{c} u[n] \\ \underline{u}_m[n-1] \end{array} \right] + \Gamma_m \left[\underline{a}_{m-1}^* \middle| 0 \right] \left[\begin{array}{c} \underline{u}_m[n] \\ u[n-m] \end{array} \right]$

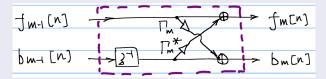
$$b_m[n] = b_{m-1}[n-1] + \Gamma_m f_{m-1}[n]$$

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Basic Lattice Structure

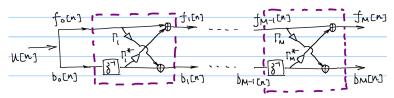
$$\begin{bmatrix} f_m[n] \\ b_m[n] \end{bmatrix} = \begin{bmatrix} 1 & \Gamma_m^* \\ \Gamma_m & 1 \end{bmatrix} \begin{bmatrix} f_{m-1}[n] \\ b_{m-1}[n-1] \end{bmatrix}, m = 1, 2, \dots, M$$

Signal Flow Graph (SFG)



Modular Structure

Recall $f_0[n] = b_0[n] = u[n]$, thus



To increase the order, we simply add more stages and reuse the earlier computations.

Using a tapped delay line implementation, we need M separate filters to generate $b_1[n], b_2[n], \ldots, b_M[n]$.

In contrast, a single lattice structure can generate $b_1[n], \ldots, b_M[n]$ as well as $f_1[n], \ldots, f_M[n]$ at the same time.

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Given from a zero-mean w.s.s. process:
$$\{u[n-1], \dots, u[n-M]\} \qquad \Rightarrow \qquad u[n]$$
 (BLP)
$$\{u[n], u[n-1], \dots, u[n-M+1]\} \qquad \Rightarrow \qquad u[n-M]$$

1. Principle of Orthogonality

$$\mathbb{E}\left[f_m[n]u^*[n-k]\right] = 0 \ (1 \le k \le m) \qquad f_m[n] \perp \underline{u}_m[n-1]$$

$$\mathbb{E}\left[b_m[n]u^*[n-k]\right] = 0 \ (0 \le k \le m-1) \qquad b_m[n] \perp \underline{u}_m[n]$$

2.
$$\mathbb{E}[f_m[n]u^*[n]] = \mathbb{E}[b_m[n]u^*[n-m]] = P_m$$

Proof: (Details)

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3. Correlations of error signals across orders:

(BLP)
$$\mathbb{E}\left[b_m[n]b_i^*[n]\right] = \begin{cases} P_m & i = m \\ 0 & i < m \text{ i.e., } b_m[n] \perp b_i[n] \end{cases}$$

(FLP)
$$\mathbb{E}[f_m[n]f_i^*[n]] = P_m \text{ for } i \leq m$$

<u>Proof</u>: (Can obtain the case i > m by conjugation)

Remark: The generation of $\{b_0[n], b_1[n], \ldots, \}$ is like a **Gram-Schmidt** orthogonalization process on $\{u[n], u[n-1], \ldots, \}$.

As a result, $\{b_i[n]\}_{i=0,1,...}$ is a new, **uncorrelated** representation of $\{u[n]\}$ containing exactly the **same information**.

4. Correlations of error signals across orders and time:

$$\mathbb{E}[f_m[n]f_i^*[n-\ell]] = \mathbb{E}[f_m[n+\ell]f_i^*[n]] = 0 \ (1 \le \ell \le m-i, i < m)$$

$$\mathbb{E}[b_m[n]b_i^*[n-\ell]] = \mathbb{E}[b_m[n+\ell]b_i^*[n]] = 0 \ (0 \le \ell \le m-i-1, i < m)$$

Proof: (Details)

5. Correlations of error signals across orders and time:

$$\mathbb{E}\left[f_m[n+m]f_i^*[n+i]\right] = \begin{cases} P_m & i=m\\ 0 & i< m \end{cases}$$

$$\mathbb{E}\left[b_m[n+m]b_i^*[n+i]\right] = P_m \qquad i \le m$$

Proof:



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6. Cross-correlations of FLP and BLP error signals:

$$\mathbb{E}\left[f_m[n]b_i^*[n]\right] = \begin{cases} \Gamma_i^* P_m & i \leq m \\ 0 & i > m \end{cases}$$

Proof:



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Joint Process Estimator: Motivation

(Readings: Haykin §3.10; Hayes §7.2.4, §9.2.8)

In (general) Wiener filtering theory, we use $\{x[n]\}$ process to estimate a desired response $\{d[n]\}$.

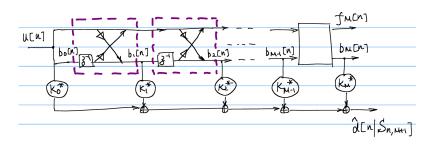
Solving the normal equation may require inverting the correlation matrix \mathbf{R}_{x} .

We now use the lattice structure to obtain a backward prediction error process $\{b_i[n]\}$ as an equivalent, uncorrelated representation of $\{u[n]\}$ that contains exactly the same information.

We then apply an optimal filter on $\{b_i[n]\}$ to estimate $\{d[n]\}$.

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Joint Process Estimator: Structure



$$\hat{d}\left[n|\mathbb{S}_{n}\right] = \underline{k}^{H}\underline{b}_{M+1}[n]$$
, where $\underline{k} = \left[k_{0}, k_{1}, \dots, k_{M}\right]^{T}$

Joint Process Estimator: Result

To determine the optimal weight to minimize MSE of estimation:

① Denote D as the $(M+1) \times (M+1)$ correlation matrix of b[n]

$$D = \mathbb{E}\left[\underline{b}[n]\underline{b}^{H}[n]\right] = \underset{\sim}{\operatorname{diag}}(P_0, P_1, \dots, P_M)$$
$$\therefore \{b_k[n]\}_{k=0}^M \text{ are uncorrelated}$$

2 Let s be the crosscorrelation vector

$$\underline{s} \triangleq [s_0, \ldots, s_M \ldots]^T = \mathbb{E}[\underline{b}[n]d^*[n]]$$

The normal equation for the optimal weight vector is:

$$D\underline{k}_{\text{opt}} = \underline{s}$$

$$\Rightarrow \underline{k}_{\text{opt}} = D^{-1}\underline{s} = \text{diag}(P_0^{-1}, P_1^{-1}, \dots, P_M^{-1})\underline{s}$$
i.e., $k_i = P_i^{-1}s_i$, $i = 0, \dots, M$

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Joint Process Estimator: Summary

The name "joint process estimation" refers to the system's structure that performs two optimal estimation jointly:

- One is a **lattice predictor** (characterized by $\Gamma_1, \ldots, \Gamma_M$) transforming a sequence of correlated samples u[n], $u[n-1], \ldots, u[n-M]$ into a sequence of uncorrelated samples $b_0[n], b_1[n], \ldots, b_M[n]$.
- The other is called a **multiple regression filter** (characterized by \underline{k}), which uses $b_0[n], \ldots, b_M[n]$ to produce an estimate of d[n].

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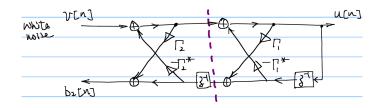
Inverse Filtering

The lattice predictor discussed just now can be viewed as an analyzer, i.e., to represent an (approximately) AR process $\{u[n]\}$ using $\{\Gamma_m\}$.

With some reconfiguration, we can obtain an inverse filter or a <u>synthesizer</u>, i.e., we can reproduce an AR process by applying white noise $\{v[n]\}$ as the input to the filter.

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A 2-stage Inverse Filtering



$$u[n] = v[n] - \Gamma_1^* u[n-1] - \Gamma_2^* (\Gamma_1 u[n-1] + u[n-2])$$

$$= v[n] - \underbrace{(\Gamma_1^* + \Gamma_1 \Gamma_2^*)}_{a_{2,1}^*} u[n-1] - \underbrace{\Gamma_2^*}_{a_{2,2}^*} u[n-2]$$

$$\therefore u[n] + a_{2,1}^* u[n-1] + a_{2,2}^* u[n-2] = v[n]$$

$$\Rightarrow \{u[n]\} \text{ is an } \mathbf{AR(2)} \text{ process.}$$

Basic Building Block for All-pole Filtering

$$\begin{cases} x_{m-1}[n] = x_m[n] - \Gamma_m^* y_{m-1}[n-1] \\ y_m[n] = \Gamma_m x_{m-1}[n] + y_{m-1}[n-1] \\ = \Gamma_m x_m[n] + (1 - |\Gamma_m|^2) y_{m-1}[n-1] \end{cases}$$

$$\Rightarrow \begin{cases} x_m[n] = x_{m-1}[n] + \Gamma_m^* y_{m-1}[n-1] \\ y_m[n] = \Gamma_m x_{m-1}[n] + y_{m-1}[n-1] \end{cases}$$

$$\therefore \left[\begin{array}{c} x_m[n] \\ y_m[n] \end{array} \right] = \left[\begin{array}{cc} 1 & \Gamma_m^* \\ \Gamma_m & 1 \end{array} \right] \left[\begin{array}{c} x_{m-1}[n] \\ y_{m-1}[n-1] \end{array} \right]$$

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All-pole Filter via Inverse Filtering

$$\begin{bmatrix} x_m[n] \\ y_m[n] \end{bmatrix} = \begin{bmatrix} 1 & \Gamma_m^* \\ \Gamma_m & 1 \end{bmatrix} \begin{bmatrix} x_{m-1}[n] \\ y_{m-1}[n-1] \end{bmatrix}$$

This gives basically the same relation as the forward lattice module:

$$x_{m-1}(n)$$
 $y_{m-1}(n)$
 $y_{m}(n) = y_{m}(n) = y_{m}(n)$

$$\Rightarrow u[n] = -a_{2,1}^* u[n-1] - a_{2,2}^* u[n-2] + v[n]$$

v[n]: white noise

